

Review Quiz #2 elec303

1

Transform of $P_X(x)$:

$$M_X(s) = \int_{-\infty}^{\infty} e^{sx} P_X(x) dx = E[e^{sx}]$$

Moment-generating function:

$$E[X^n] = \left. \frac{d^n}{ds} M_X(s) \right|_{s=0}$$

the n^{th} moment of X

Sums of Independent RVs

$Z = X + Y$, (X, Y) are independent.

$$\begin{aligned} P_Z(z) &= P(X + Y = z) \\ &= \int_{-\infty}^{\infty} P_{X,Z}(x, z) dz = \int_{-\infty}^{\infty} P_X(x) P_Y(z - x) dx \end{aligned}$$

convolution of $P_X(x)$ w/ $P_Y(y)$.

Alternatively,

2

$$Z = X + Y.$$

Compute the transform of z .

$$\begin{aligned} M_Z(s) E[e^{sZ}] &= E[e^{s(X+Y)}] = E[e^{sX} e^{sY}] \\ &= E[e^{sX}] E[e^{sY}] \\ &= M_X(s) \cdot M_Y(s) \end{aligned}$$

$P_Z(z) =$ inverse transform of $M_X(s) \cdot M_Y(s)$.

(Section 4.5) Sum of a random number of independ. RVs

$$Y = X_1 + \dots + X_n$$

where n is random
[takes on non-neg. integer values!]

&

X_i 's are i.i.d
(independent & identical distributed)

$$P_Y(y) = ?$$

Let $E[X_i] = \mu$
 $\text{VAR}(X_i) = \sigma_x^2$ for all i . 3

The random variable

$Y = X_1 + \dots + X_n$ is independent of N
and is therefore independent of $\{N=n\}$.

Hence,

$$\begin{aligned} E[Y | N=n] &= E[X_1 + \dots + X_n | N=n] \\ &= E[X_1 + \dots + X_n] \\ &= n E[X] = n\mu. \end{aligned}$$

This is true for any non-negative integer n , so

$$E[Y | N] = N E[X].$$

Using the law of iterated expectations,
we obtain:

$$E[Y] = E[E[Y | N]] = E[N E[X]] = \underline{E[N] E[X]}.$$

$$\begin{aligned} \text{VAR}(Y | N=n) &= \text{VAR}(X_1 + \dots + X_N | N=n) & 4 \\ &= \text{VAR}(X_1 + \dots + X_n) \\ &= n \cdot \text{VAR}(X) = \underline{n \sigma_x^2}. \end{aligned}$$

$$\begin{aligned} \Rightarrow \text{VAR}(Y) &= E[\text{VAR}(Y|N)] + \text{VAR}(E[Y|N]) \\ &= E[N \text{VAR}(X)] + \text{VAR}(N E[X]) \\ &= E[N] \text{VAR}(X) + (E[X])^2 \text{VAR}(N). \end{aligned}$$

Thus to find the transform of Y ,

$$\begin{aligned} M_Y(s) &= E[e^{sY}] = E[E[e^{sY} | N]] \overset{E[e^{sX_1} \cdot e^{sX_2} \dots e^{sX_N}]}{=} E[(M_X(s))^N] \\ &= E[(M_X(s))^N] \end{aligned}$$

Using

$$(M_X(s))^n = e^{\log(M_X(s))^n} = \cancel{e^{n \log(M_X(s))}}$$

$$= \underline{\exp(n \log(M_X(s)))}$$

We have that :

5

$$M_Y(s) = E_N[(M_X(s))^N]$$

$\hookrightarrow N$ is discrete & non-neg.

$$= \sum_{n=0}^{\infty} e^{n \log(M_X(s))} \cdot P_N(n)$$

and

$$M_N(s) = E[e^{sN}] = \sum e^{sn} P_N(n)$$

\Downarrow

implies that,

$$M_Y(s) = M_N(\log(M_X(s))) \longleftrightarrow P_Y(y)$$

inverse transf.

Summary

for $Y = X_1 + \dots + X_N$, where N is random [integer & non-negative] & X_i 's are i.i.d.

(1) $E[Y] = E[N] E[X]$.

(2) $\text{VAR}(Y) = E[N] \text{VAR}(X) + (E[X])^2 \text{VAR}(N)$.

(3) $M_Y(s) = M_N(\log M_X(s))$.

COVARIANCE & CORRELATION

6

(1) COVARIANCE

$$\text{cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$$

$$= E[XY] - 2E[X]E[Y] + E[X]E[Y]$$

$$= E[XY] - E[X]E[Y].$$

if X, Y are independent?

$$E[XY] = E[X]E[Y]$$

$$\Rightarrow \text{cov}(X, Y) = 0 !$$

(2) CORRELATION COEFFICIENT

$$\rho(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{\text{VAR}(X)\text{VAR}(Y)}}$$

Variance of a sum of RANDOM VARIABLES

7

$$\text{VAR}(X_1 + X_2) = \text{VAR}(X_1) + \text{VAR}(X_2) \\ + 2 \text{COV}(X_1, X_2)$$

[see notes posted online]

* define $\tilde{X}_i = X_i - E[X_i]$ *

$$\text{VAR}\left(\sum_{i=1}^n X_i\right) = E\left[\left(\sum_{i=1}^n \tilde{X}_i\right)^2\right]$$

$$= E\left[\sum_{i=1}^n \sum_{j=1}^n \tilde{X}_i \tilde{X}_j\right]$$

$$= \sum_{i=1}^n \sum_{j=1}^n E[\tilde{X}_i \tilde{X}_j]$$

$$= \sum_{i=1}^n E[\tilde{X}_i^2] + \sum_{\{(i,j), i \neq j\}} E[\tilde{X}_i \tilde{X}_j]$$

$$= \sum_{i=1}^n \text{VAR}(X_i) + \sum_{\substack{i,j \\ i \neq j}} \text{COV}(X_i, X_j)$$

Problem 27 (Pg. 253)

8

We toss a biased coin ($P(\text{head}) = q$) n times, where q is a value of a RV Q w/ mean = μ , $\text{VAR} = \sigma^2$.

Let X_i be a Bernoulli RV that models the outcome of the i^{th} toss ($X_i = 1$ if i^{th} toss is a head).

Assume that X_1, \dots, X_n are conditionally independent, given $Q = q$. Let $X = \sum_{i=1}^n X_i$

be the # of heads in n tosses.

(a) Use the law of iterated expectations to find $E[X_i]$ and $E[X]$.

$$\otimes E[X_i | Q] = Q$$

[expectation of a Bernoulli w/ parameter q]

$$\begin{aligned} E[X_i] &= E[E[X_i | Q]] \\ &= E[Q] = \underline{\mu}. \end{aligned}$$

$$\begin{aligned} \downarrow E[X_i | Q] &= \sum_{x_i=0}^1 x_i P_{X_i}(x_i) \\ &= (0)(1-q) + (1)q \\ &= \underline{q}. \end{aligned}$$

Since $X = X_1 + \dots + X_n$

9

$$\begin{aligned} E[X] &= E[X_1] + \dots + E[X_n] \\ &= \underline{n\mu}. \end{aligned}$$

(b) Find $\text{cov}(X_i, X_j)$. Are X_1, \dots, X_n independent?

⊗ For $i \neq j$, (X_i, X_j) assumed to be condit. independent!!

$$\begin{aligned} E[X_i X_j | \mathcal{Q}] &= E[X_i | \mathcal{Q}] \cdot E[X_j | \mathcal{Q}] \\ &= \underline{\mathcal{Q}^2}. \end{aligned}$$

$$\begin{aligned} E[X_i X_j] &= E_{\mathcal{Q}}[E[X_i X_j | \mathcal{Q}]] \\ &= \underline{E[\mathcal{Q}^2]}. \end{aligned}$$

Thus,

$$\begin{aligned} \text{cov}(X_i, X_j) &= E[X_i X_j] - E[X_i]E[X_j] \\ &= E[\mathcal{Q}^2] - \mu^2 = \underline{\sigma^2}. \end{aligned}$$

Since $\text{cov}(X_i, X_j) \neq 0 \Rightarrow X_i, X_j$ are NOT independent!!

For $i=j$ $\rightarrow X_i^2 = X_i$

10

$$\begin{aligned}\text{VAR}(X_i) &= E[X_i^2] - (E[X_i])^2 \\ &= E[X_i] - (E[X_i])^2 \\ &= \underline{\mu - \mu^2}.\end{aligned}$$

(c) Use the law of total variance to find $\text{VAR}(X)$.
Verify your answer using the cov in part (b).

$$\begin{aligned}\text{VAR}(X) &= E[\text{VAR}(X|\mathcal{Q})] + \text{VAR}(E[X|\mathcal{Q}]) \\ &= E[\text{VAR}(X_1 + \dots + X_n | \mathcal{Q})] + \text{VAR}(E[X_1 + \dots + X_n | \mathcal{Q}]) \\ &= \underbrace{E[n\mathcal{Q}(1-\mathcal{Q})]}_{\text{condit. independ.}} + \underbrace{\text{VAR}(n\mathcal{Q})}_{\text{VAR}(aX) = a^2 \text{VAR}(X)} \\ &= n E[\mathcal{Q} - \mathcal{Q}^2] + n^2 \text{VAR}(\mathcal{Q}) \\ &= n(\mu - \overset{\uparrow}{\mu^2} - \sigma^2) + n^2 \sigma^2 = \underline{n(\mu - \mu^2) + n(n-1)\sigma^2}.\end{aligned}$$

To verify the results
from part (b),

//

$$\text{VAR}(X) = \text{VAR}(X_1 + \dots + X_n)$$

$$= \sum_{i=1}^n \text{VAR}(X_i) + \sum_{i \neq j} \text{COV}(X_i, X_j)$$

$$= n \text{VAR}(X_i) + n(n-1) \text{COV}(X_1, X_2)$$

$$= n(\mu - \mu^2) + n(n-1)\sigma^2$$

Same as before !!

Review -

12

Normal Approximation Based on CLT

$$\text{Let } S_n = X_1 + \dots + X_n,$$

where X_i 's are i.i.d w/ mean = μ
VAR = σ^2 .

$$P(S_n \leq c)$$

$$= P\left(\frac{S_n - n\mu}{\sigma\sqrt{n}} \leq \frac{c - n\mu}{\sigma\sqrt{n}}\right)$$

$$= P(Z_n \leq \frac{c - n\mu}{\sigma\sqrt{n}})$$

$$\approx \Phi\left(\frac{c - n\mu}{\sigma\sqrt{n}}\right)$$

↳ CDF of a STANDARD normal w/ mean = 0
VAR = 1.

A note on Gaussian CDFs. 13

$$Y \sim N(\mu, \sigma^2), \quad X \sim \underbrace{N(0, 1)}_{\text{STANDARD NORMAL}}$$

$$\underline{Y = \sigma X + \mu} \quad \cdot \quad (\text{Property of Gaussians})$$

Thus for

$$\Phi(c) = \int_{-\infty}^c e^{-x^2/2} dx = P(X \leq c),$$

We can express the CDF of an arbitrary Gaussian in terms of the

$$P(Y \leq c) = \left. \begin{array}{l} \text{STANDARD} \\ \text{NORMAL CDF} \end{array} \right\}$$

$$= P(\sigma X + \mu \leq c) = P\left(X \leq \frac{c - \mu}{\sigma}\right)$$

$$= \Phi\left(\frac{c - \mu}{\sigma}\right).$$